

9231_41_Summer_2020_Q1

Solution

To determine whether the grades awarded are independent of the presence of background music, we perform a **Chi-square test for independence**.

1. Hypotheses and Significance Level

- **Null Hypothesis** (H_0): Grade awarded and background music are independent.
- **Alternative Hypothesis** (H_1): Grade awarded and background music are not independent.
- Significance level: $\alpha = 0.10$.

2. Observed Frequencies and Totals The observed frequencies (O_{ij}) are provided in the contingency table. We first calculate the row totals, column totals, and the grand total (N).

| | Grade A | Grade B | Grade C | Row Total |
|--------------|---------|---------|---------|-----------|
| Music | 49 | 51 | 40 | 140 |
| Silence | 93 | 68 | 49 | 210 |
| Column Total | 142 | 119 | 89 | 350 |

3. Expected Frequencies Under the assumption of independence, the **expected frequency** (E_{ij}) for each cell is calculated as:

$$E_{ij} = \frac{(\text{Row Total}) \times (\text{Column Total})}{\text{Grand Total}}$$

- For Music and Grade A: $E_{11} = \frac{140 \times 142}{350} = 56.8$
- For Music and Grade B: $E_{12} = \frac{140 \times 119}{350} = 47.6$
- For Music and Grade C: $E_{13} = \frac{140 \times 89}{350} = 35.6$
- For Silence and Grade A: $E_{21} = \frac{210 \times 142}{350} = 85.2$
- For Silence and Grade B: $E_{22} = \frac{210 \times 119}{350} = 71.4$
- For Silence and Grade C: $E_{23} = \frac{210 \times 89}{350} = 53.4$

4. Test Statistic Calculation The **Chi-square test statistic** (χ^2) is given by:

$$\chi^2 = \sum \frac{(O_{ij} - E_{ij})^2}{E_{ij}}$$

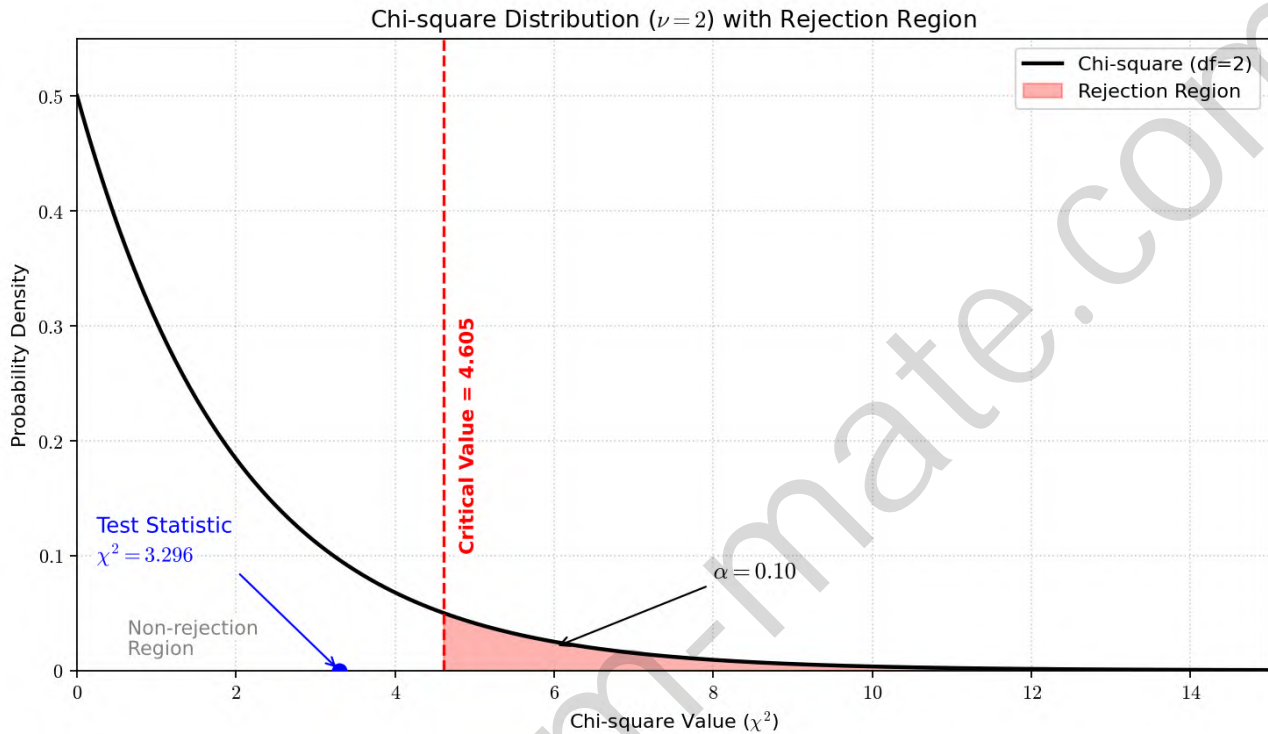
$$\begin{aligned} \chi^2 &= \frac{(49 - 56.8)^2}{56.8} + \frac{(51 - 47.6)^2}{47.6} + \frac{(40 - 35.6)^2}{35.6} + \frac{(93 - 85.2)^2}{85.2} + \frac{(68 - 71.4)^2}{71.4} + \frac{(49 - 53.4)^2}{53.4} \\ &= \frac{(-7.8)^2}{56.8} + \frac{3.4^2}{47.6} + \frac{4.4^2}{35.6} + \frac{7.8^2}{85.2} + \frac{(-3.4)^2}{71.4} + \frac{(-4.4)^2}{53.4} \\ &\approx 1.0711 + 0.2429 + 0.5438 + 0.7141 + 0.1619 + 0.3625 \\ &\approx 3.2963 \end{aligned}$$

5. Degrees of Freedom and Critical Value The **degrees of freedom** (df) for a contingency table with r rows and c columns is:

$$df = (r - 1)(c - 1) = (2 - 1)(3 - 1) = 2$$

Using the **Chi-square distribution** table for $df = 2$ at $\alpha = 0.10$:

$$\chi_{0.10,2}^2 = 4.605$$



6. Conclusion Since the calculated test statistic $\chi^2 \approx 3.296$ is less than the critical value 4.605, we fail to reject the null hypothesis H_0 .

There is insufficient evidence at the 10% significance level to suggest that the grades awarded are dependent on whether background music is playing. We conclude that the grades and the presence of music are independent.

Fail to reject H_0 ; grades and music are independent at the 10% level.

9231_41_Summer_2020_Q2

Solution

The problem asks for a **Wilcoxon signed-rank test** to evaluate a claim about the median of a population based on a sample of 10 observations.

1. Hypotheses and Significance Level Let m be the population median time to complete the task.

- **Null Hypothesis** $H_0 : m = 6.4$ ms
- **Alternative Hypothesis** $H_1 : m \neq 6.4$ ms (two-tailed test)
- **Significance Level** $\alpha = 0.05$

2. Calculation of Differences and Ranks We calculate the difference $d_i = x_i - 6.4$ for each observation x_i , then rank the absolute differences $|d_i|$.

| x_i | $d_i = x_i - 6.4$ | $ d_i $ | Rank | Signed Rank |
|-------|-------------------|---------|------|-------------|
| 6.44 | +0.04 | 0.04 | 1 | +1 |
| 6.16 | -0.24 | 0.24 | 7 | -7 |
| 5.62 | -0.78 | 0.78 | 10 | -10 |
| 5.82 | -0.58 | 0.58 | 9 | -9 |
| 6.51 | +0.11 | 0.11 | 4 | +4 |
| 6.62 | +0.22 | 0.22 | 6 | +6 |
| 6.19 | -0.21 | 0.21 | 5 | -5 |
| 6.42 | +0.02 | 0.02 | 2 | +2 |
| 6.34 | -0.06 | 0.06 | 3 | -3 |
| 6.28 | -0.12 | 0.12 | 8 | -8 |

3. Test Statistic Calculation The test statistic W is the smaller of the sum of positive ranks (W^+) and the sum of negative ranks (W^-).

- Sum of positive ranks:

$$W^+ = 1 + 4 + 6 + 2 = 13$$

- Sum of negative ranks:

$$W^- = 7 + 10 + 9 + 5 + 3 + 8 = 42$$

- Test statistic:

$$T = \min(W^+, W^-) = 13$$

4. Critical Value and Conclusion For a two-tailed test with $n = 10$ at $\alpha = 0.05$:

- From the **Wilcoxon signed-rank table**, the critical value T_{crit} is 8.
- The decision rule for this test is to reject H_0 if $T \leq T_{\text{crit}}$.
- Since $13 > 8$, we fail to reject the null hypothesis.

There is insufficient evidence at the 5% significance level to suggest that the median time to complete the task is not 6.4 ms. The claim is supported by the data.

5. Underlying Assumption The **Wilcoxon signed-rank test** requires that the population distribution from which the sample is drawn is **symmetrical** about its median.

Final Answers:

(a)

$$H_0 : m = 6.4$$

$$H_1 : m \neq 6.4$$

$$\text{Sum of positive ranks } (W^+) = 13$$

$$\text{Sum of negative ranks } (W^-) = 42$$

$$\text{Test statistic } T = 13$$

$$\text{Critical value } (n = 10, \text{ two-tailed, } 5\%) = 8$$

Since $13 > 8$, do not reject H_0 . There is no significant evidence to reject the claim that the median is 6.4 ms.

(b) The distribution of the population must be **symmetrical**.

9231_41_Summer_2020_Q3

Solution

The continuous random variable X is defined by the **probability density function** (PDF) $f(x)$:

$$f(x) = \begin{cases} \frac{3}{16}(2 - \sqrt{x}) & 0 \leq x < 1 \\ \frac{3}{16\sqrt{x}} & 1 \leq x \leq 9 \\ 0 & \text{otherwise} \end{cases}$$

1. Calculation of the Expected Value $E(X)$

The **expected value** of a continuous random variable is defined by the integral:

$$E(X) = \int_{-\infty}^{\infty} xf(x) dx$$

Given the piecewise nature of $f(x)$, we split the integral into two non-zero regions:

- **Region 1:** $0 \leq x < 1$

$$\begin{aligned} I_1 &= \int_0^1 x \cdot \frac{3}{16}(2 - \sqrt{x}) dx \\ &= \frac{3}{16} \int_0^1 (2x - x^{3/2}) dx \\ &= \frac{3}{16} \left[x^2 - \frac{2}{5}x^{5/2} \right]_0^1 \\ &= \frac{3}{16} \left(1 - \frac{2}{5} \right) = \frac{3}{16} \cdot \frac{3}{5} = \frac{9}{80} \end{aligned}$$

- **Region 2:** $1 \leq x \leq 9$

$$\begin{aligned} I_2 &= \int_1^9 x \cdot \frac{3}{16\sqrt{x}} dx \\ &= \frac{3}{16} \int_1^9 x^{1/2} dx \\ &= \frac{3}{16} \left[\frac{2}{3}x^{3/2} \right]_1^9 \\ &= \frac{1}{8} (9^{3/2} - 1^{3/2}) \\ &= \frac{1}{8} (27 - 1) = \frac{26}{8} = \frac{13}{4} \end{aligned}$$

- **Total Expected Value:**

$$\begin{aligned}
 E(X) &= I_1 + I_2 \\
 &= \frac{9}{80} + \frac{13}{4} \\
 &= \frac{9}{80} + \frac{260}{80} = \frac{269}{80} = 3.3625
 \end{aligned}$$

2. Probability Density Function of $Y = \sqrt{X}$

To find the PDF of Y , we use the **method of transformations**. Let $g(x) = \sqrt{x}$. Since $g(x)$ is strictly increasing for $x \geq 0$, we have $X = Y^2$. The PDF of Y is given by:

$$f_Y(y) = f_X(x) \left| \frac{dx}{dy} \right|$$

where $x = y^2$ and $\frac{dx}{dy} = 2y$.

- **Determine the range of Y :**

- For $0 \leq x < 1$, $0 \leq y < 1$.
- For $1 \leq x \leq 9$, $1 \leq y \leq 3$.

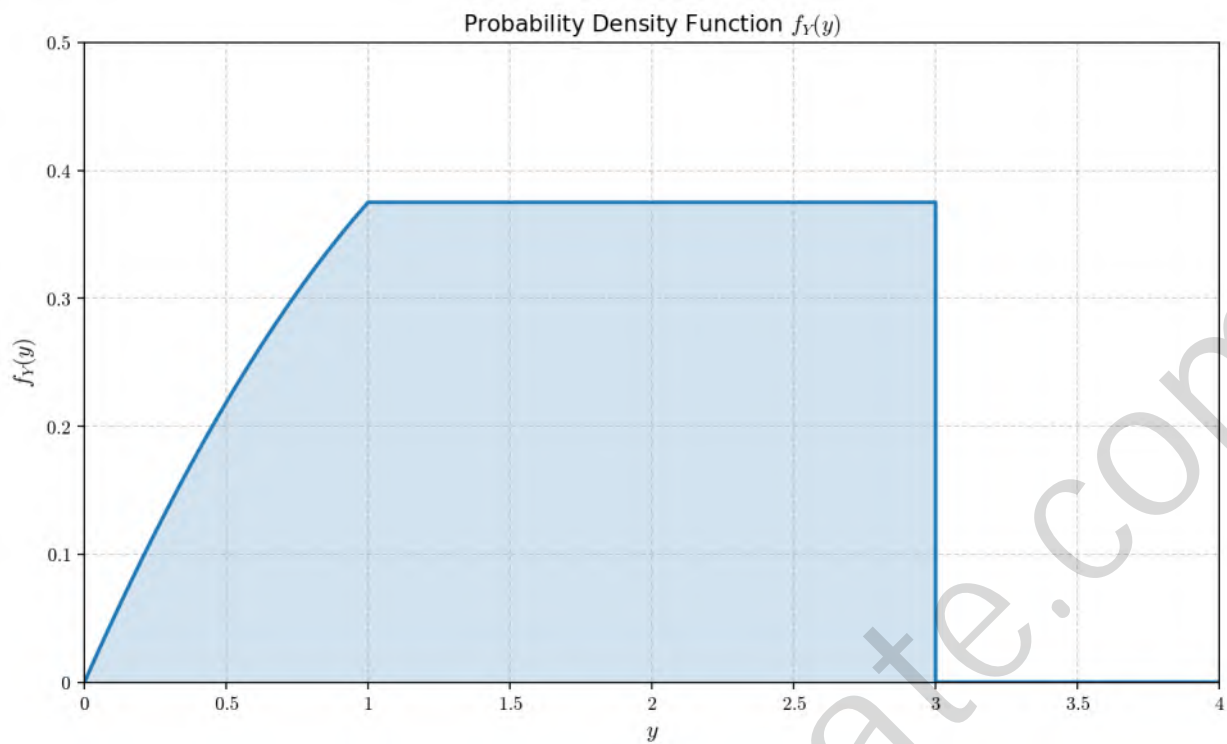
- **Transforming the PDF:**

- For $0 \leq y < 1$:

$$\begin{aligned}
 f_Y(y) &= f_X(y^2) \cdot 2y \\
 &= \frac{3}{16}(2 - \sqrt{y^2}) \cdot 2y \\
 &= \frac{3}{16}(2 - y) \cdot 2y = \frac{3y}{8}(2 - y)
 \end{aligned}$$

- For $1 \leq y \leq 3$:

$$\begin{aligned}
 f_Y(y) &= f_X(y^2) \cdot 2y \\
 &= \frac{3}{16\sqrt{y^2}} \cdot 2y \\
 &= \frac{3}{16y} \cdot 2y = \frac{3}{8}
 \end{aligned}$$



The final PDF for Y is:

$$f_Y(y) = \begin{cases} \frac{3y}{8}(2-y) & 0 \leq y < 1 \\ \frac{3}{8} & 1 \leq y \leq 3 \\ 0 & \text{otherwise} \end{cases}$$

Final Answers:

(a) 3.3625

(b) $f_Y(y) = \begin{cases} \frac{3y}{8}(2-y) & 0 \leq y < 1 \\ \frac{3}{8} & 1 \leq y \leq 3 \\ 0 & \text{otherwise} \end{cases}$

9231_41_Summer_2020_Q4

Solution

To test the manager's claim that there is no difference between the mean volume of coffee filled by machine X and machine Y , we perform a **two-sample t-test** (or z-test, given the large sample sizes).

1. State the Hypotheses Let μ_x and μ_y be the population mean volumes for machines X and Y , respectively.

- **Null Hypothesis** (H_0): $\mu_x = \mu_y$ (or $\mu_x - \mu_y = 0$)
- **Alternative Hypothesis** (H_1): $\mu_x \neq \mu_y$ (Two-tailed test)
- Significance level: $\alpha = 0.10$

2. Calculate Sample Statistics For Machine X ($n_x = 50$):

- Sample mean: $\bar{x} = \frac{\sum x}{n_x} = \frac{15.2}{50} = 0.304$
- Unbiased estimate of variance (s_x^2):

$$\begin{aligned} s_x^2 &= \frac{1}{n_x - 1} \left(\sum x^2 - \frac{(\sum x)^2}{n_x} \right) \\ &= \frac{1}{49} \left(5.1 - \frac{15.2^2}{50} \right) \\ &= \frac{1}{49} (5.1 - 4.6208) \\ &= \frac{0.4792}{49} \approx 0.00977959 \end{aligned}$$

For Machine Y ($n_y = 40$):

- Sample mean: $\bar{y} = \frac{\sum y}{n_y} = \frac{13.4}{40} = 0.335$
- Unbiased estimate of variance (s_y^2):

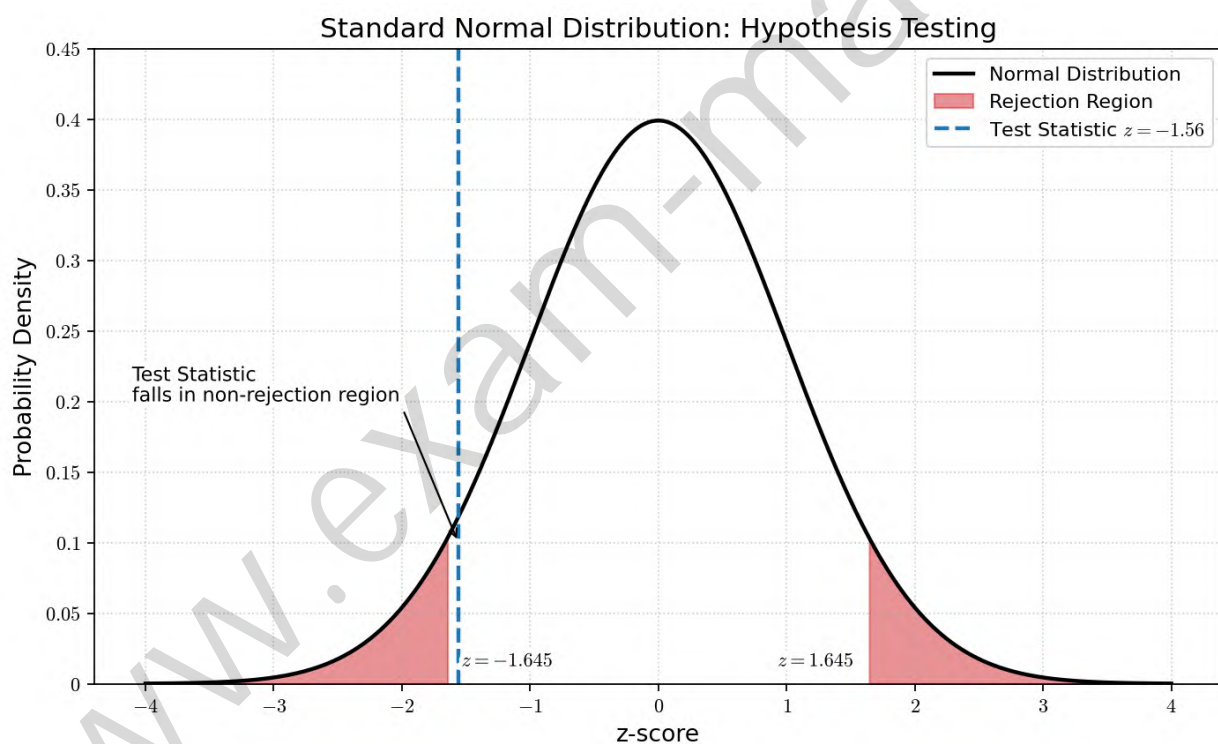
$$\begin{aligned} s_y^2 &= \frac{1}{n_y - 1} \left(\sum y^2 - \frac{(\sum y)^2}{n_y} \right) \\ &= \frac{1}{39} \left(4.8 - \frac{13.4^2}{40} \right) \\ &= \frac{1}{39} (4.8 - 4.489) \\ &= \frac{0.311}{39} \approx 0.00797436 \end{aligned}$$

3. Calculate the Test Statistic Since the sample sizes are large ($n_x, n_y \geq 30$), we use the **Central Limit Theorem** to justify the use of the **standard normal distribution** (Z). The test statistic is:

$$\begin{aligned}
 z &= \frac{(\bar{x} - \bar{y}) - 0}{\sqrt{\frac{s_x^2}{n_x} + \frac{s_y^2}{n_y}}} \\
 &= \frac{0.304 - 0.335}{\sqrt{\frac{0.00977959}{50} + \frac{0.00797436}{40}}} \\
 &= \frac{-0.031}{\sqrt{0.00019559 + 0.00019936}} \\
 &= \frac{-0.031}{\sqrt{0.00039495}} \\
 &\approx \frac{-0.031}{0.019873} \\
 &\approx -1.5599
 \end{aligned}$$

4. Determine the Critical Value and Conclusion For a two-tailed test at the 10% significance level ($\alpha = 0.10$), the critical values are obtained from the standard normal table:

- $P(Z > z_c) = 0.05 \implies z_c = 1.645$
- The **rejection region** is $|z| > 1.645$.



Comparing the calculated value to the critical value:

$$|-1.560| = 1.560 < 1.645$$

Since the test statistic does not fall in the rejection region, we fail to reject the null hypothesis. There is insufficient evidence at the 10% significance level to suggest a difference between the mean volumes of coffee filled by the two machines. The manager's claim is supported by the data.

The calculated test statistic is $z \approx -1.56$. Since $|-1.56| < 1.645$, we **accept** H_0 . There is no significant evidence of a difference in means.

9231_41_Summer_2020_Q5

Solution

1. Descriptive Statistics

First, we calculate the sample mean and sample standard deviation from the given data set of $n = 8$ distances: $\{19.8, 22.1, 24.4, 21.5, 20.8, 26.3, 23.7, 25.0\}$.

- The **sample mean** (\bar{x}) is:

$$\begin{aligned}\bar{x} &= \frac{\sum x_i}{n} \\ &= \frac{19.8 + 22.1 + 24.4 + 21.5 + 20.8 + 26.3 + 23.7 + 25.0}{8} \\ &= \frac{183.6}{8} \\ &= 22.95 \text{ m}\end{aligned}$$

- The **unbiased sample variance** (s^2) is:

$$\begin{aligned}s^2 &= \frac{\sum (x_i - \bar{x})^2}{n - 1} \\ &= \frac{(19.8 - 22.95)^2 + \dots + (25.0 - 22.95)^2}{7} \\ &= \frac{35.42}{7} \\ &= 5.06\end{aligned}$$

- The **sample standard deviation** (s) is:

$$s = \sqrt{5.06} \approx 2.24944 \text{ m}$$

2. Hypothesis Test (Part a)

We perform a one-sample **t-test** because the population variance is unknown and the sample size is small ($n < 30$).

- **Hypotheses:**

- **Null Hypothesis** $H_0 : \mu = 22.0$
- **Alternative Hypothesis** $H_1 : \mu > 22.0$ (one-tailed test)

- **Test Statistic:**

$$\begin{aligned}t &= \frac{\bar{x} - \mu_0}{s/\sqrt{n}} \\ &= \frac{22.95 - 22.0}{2.24944/\sqrt{8}} \\ &= \frac{0.95}{0.7953} \\ &\approx 1.1945\end{aligned}$$

- **Critical Value:** For a significance level $\alpha = 0.05$ and **degrees of freedom** $df = n - 1 = 7$, the critical value from the **Student's t-distribution** table is:

$$t_{0.05,7} = 1.895$$

- **Conclusion:** Since the calculated t -value (1.1945) is less than the critical value (1.895), we fail to reject the null hypothesis. There is insufficient evidence at the 5% significance level to suggest that the population mean distance is more than 22.0 metres.

3. 95% Confidence Interval (Part b)

The formula for the **confidence interval** of the mean is:

$$\bar{x} \pm t_{\alpha/2,df} \cdot \left(\frac{s}{\sqrt{n}} \right)$$

- For a 95% confidence level, $\alpha = 0.05$, so $\alpha/2 = 0.025$.
- The critical value for $df = 7$ is $t_{0.025,7} = 2.365$.

- **Calculation:**

$$\begin{aligned} \text{CI} &= 22.95 \pm 2.365 \cdot \left(\frac{2.24944}{\sqrt{8}} \right) \\ &= 22.95 \pm 2.365 \cdot (0.7953) \\ &= 22.95 \pm 1.8809 \end{aligned}$$

- **Lower Bound:** $22.95 - 1.8809 = 21.0691$
- **Upper Bound:** $22.95 + 1.8809 = 24.8309$

Rounding to three significant figures:

$$\boxed{[21.1, 24.8]}$$

9231_41_Summer_2020_Q6

Solution

1. Probability Generating Function of X

The random variable X represents the number of red balls selected when 2 balls are chosen from a bag containing 4 red and 6 blue balls without replacement. This follows a **hypergeometric distribution**. The possible values for X are $\{0, 1, 2\}$.

- The total number of ways to choose 2 balls from 10 is $\binom{10}{2} = 45$.
- $P(X = 0) = \frac{\binom{4}{0}\binom{6}{2}}{45} = \frac{1 \times 15}{45} = \frac{1}{3}$
- $P(X = 1) = \frac{\binom{4}{1}\binom{6}{1}}{45} = \frac{4 \times 6}{45} = \frac{24}{45} = \frac{8}{15}$
- $P(X = 2) = \frac{\binom{4}{2}\binom{6}{0}}{45} = \frac{6 \times 1}{45} = \frac{6}{45} = \frac{2}{15}$

The **probability generating function** (PGF) is defined as $G_X(t) = \sum P(X = x)t^x$:

$$\begin{aligned} G_X(t) &= \frac{1}{3}t^0 + \frac{8}{15}t^1 + \frac{2}{15}t^2 \\ &= \frac{5}{15} + \frac{8}{15}t + \frac{2}{15}t^2 \end{aligned}$$

$$G_X(t) = \frac{1}{15}(5 + 8t + 2t^2)$$

2. Probability Generating Function of Y

Y is the sum of two independent Bernoulli trials (coin tosses). Let Y_1 be the result of the first coin ($P(\text{Head}) = 2/3$) and Y_2 be the result of the second coin ($P(\text{Head}) = p$). The PGF of a sum of independent variables is the product of their individual PGFs:

- $G_{Y_1}(t) = \frac{1}{3} + \frac{2}{3}t$
- $G_{Y_2}(t) = (1 - p) + pt$
- $G_Y(t) = G_{Y_1}(t)G_{Y_2}(t) = \left(\frac{1}{3} + \frac{2}{3}t\right)((1 - p) + pt)$

Expanding $G_Y(t)$:

$$G_Y(t) = \frac{1}{3}(1 - p) + \left[\frac{1}{3}p + \frac{2}{3}(1 - p)\right]t + \frac{2}{3}pt^2$$

The coefficient of t is given as $7/12$:

$$\frac{1}{3}p + \frac{2}{3} - \frac{2}{3}p = \frac{7}{12}$$

$$\frac{2}{3} - \frac{1}{3}p = \frac{7}{12}$$

$$\frac{1}{3}p = \frac{8}{12} - \frac{7}{12} = \frac{1}{12}$$

$$p = \frac{1}{4}$$

Substituting $p = 1/4$ back into $G_Y(t)$:

$$\begin{aligned} G_Y(t) &= \left(\frac{1}{3} + \frac{2}{3}t\right)\left(\frac{3}{4} + \frac{1}{4}t\right) \\ &= \frac{3}{12} + \frac{1}{12}t + \frac{6}{12}t + \frac{2}{12}t^2 \\ &= \frac{1}{4} + \frac{7}{12}t + \frac{1}{6}t^2 \end{aligned}$$

$$G_Y(t) = \frac{1}{12}(3 + 7t + 2t^2)$$

3. Probability Generating Function of Z

Since $Z = X + Y$ and X, Y are independent, $G_Z(t) = G_X(t) \cdot G_Y(t)$:

$$\begin{aligned} G_Z(t) &= \left[\frac{1}{15}(5 + 8t + 2t^2)\right]\left[\frac{1}{12}(3 + 7t + 2t^2)\right] \\ &= \frac{1}{180}(5 + 8t + 2t^2)(3 + 7t + 2t^2) \\ &= \frac{1}{180}[5(3 + 7t + 2t^2) + 8t(3 + 7t + 2t^2) + 2t^2(3 + 7t + 2t^2)] \\ &= \frac{1}{180}[15 + 35t + 10t^2 + 24t + 56t^2 + 16t^3 + 6t^2 + 14t^3 + 4t^4] \\ &= \frac{1}{180}(15 + 59t + 72t^2 + 30t^3 + 4t^4) \end{aligned}$$

$$G_Z(t) = \frac{1}{12} + \frac{59}{180}t + \frac{2}{5}t^2 + \frac{1}{6}t^3 + \frac{1}{45}t^4$$

4. Expected Value of Z

The **expected value** can be found using the property $E(Z) = G_Z'(1)$. Alternatively, by the **linearity of expectation**, $E(Z) = E(X) + E(Y)$.

- For X : $E(X) = G_{X'}(1)$.

$$G_{X'}(t) = \frac{1}{15}(8 + 4t)$$

$$E(X) = G_{X'}(1) = \frac{12}{15} = 0.80$$

- For Y : $E(Y) = G_{Y'}(1)$.

$$G_{Y'}(t) = \frac{1}{12}(7 + 4t)$$

$$E(Y) = G_{Y'}(1) = \frac{11}{12} \approx 0.9167$$

- For Z :

$$\begin{aligned} E(Z) &= \frac{12}{15} + \frac{11}{12} \\ &= \frac{48}{60} + \frac{55}{60} \\ &= \frac{103}{60} \end{aligned}$$

$$E(Z) = \frac{103}{60} \approx 1.717$$

www.exam-mate.com